

THE COVARIANCE MATRIX OF NORMAL ORDER STATISTICS

BY

C. S. DAVIS and M. A. STEPHENS

TECHNICAL REPORT NO. 14

FEBRUARY 21, 1978

PREPARED UNDER GRANT

DAAG29-77-G-0031

FOR THE U.S. ARMY RESEARCH OFFICE

Reproduction in Whole or in Part is Permitted
for any purpose of the United States Government

Approved for public release; distribution unlimited.

DEPARTMENT OF STATISTICS

STANFORD UNIVERSITY

STANFORD, CALIFORNIA



THE COVARIANCE MATRIX OF NORMAL ORDER STATISTICS

By

C.S. Davis and M.A. Stephens

TECHNICAL REPORT NO. 14

FEBRUARY 21, 1978

Prepared under Grant DAAG29-77-G-0031

For the U.S. Army Research Office

Herbert Solomon, Project Director

Approved for public release; distribution unlimited.

DEPARTMENT OF STATISTICS
STANFORD UNIVERSITY
STANFORD, CALIFORNIA

Partially supported under Office of Naval Research Contract N00014-76-C-0475
(NR-042-267) and issued as Technical Report No. 254.

THE COVARIANCE MATRIX OF NORMAL ORDER STATISTICS

C.S. Davis and M.A. Stephens

McMaster University
Hamilton, Ontario, Canada

ABSTRACT

An approximation is given to calculate V , the covariance matrix for normal order statistics. The approximation gives considerable improvement over previous approximations, and the computing algorithm is available from the authors.

1. INTRODUCTION

Many statistical methods involve order statistics, and for a proper study of these methods the covariance matrix V of a sample of order statistics is needed. For a few important distributions (e.g., the uniform and exponential), the entries V_{ij} can be expressed in closed form and can be calculated easily; but for most parent populations each V_{ij} involves a double integral, so that accurate tabulation is difficult and expensive. In particular, for the normal population, V has so far been published only for samples of size $n \leq 20$, (see e.g., Sarhan and Greenberg, 1956; Owen, 1962). The need for good tables of V , for

many populations, was pointed out by Hastings et al. (1947) and the magnitude of the problem of exact calculation was also stressed; subsequently, series expansions for V_{ij} have been given by Plackett (1958) and by David and Johnson (1954). Saw (1960) compared these expansions and concluded that although Plackett's series converges a little faster for a normal population, there were computational advantages in the David-Johnson method. The David-Johnson formulae give V_{ij} up to terms in $(n+2)^{-3}$.

In this paper we are concerned with V for normal order statistics. We give a technique by which one can obtain an excellent approximation for V , by starting with the values given by the David-Johnson formulae and modifying them by use of certain identities and specially tabulated values for normal order statistics.

Suppose $X_{(1)}, X_{(2)}, \dots, X_{(n)}$ are the order statistics (in ascending order) of a sample of size n from a normal distribution with mean 0 and variance 1; let $m_i = E(X_{(i)})$, where E stands for expectation, and let V have entries $V_{ij} = E(X_{(i)} - m_i)(X_{(j)} - m_j)$. Three useful identities are:

$$\text{For any } i, \sum_{j=1}^n V_{ij} = 1 \quad ; \quad (1)$$

$$E(X_{(1)}^2) = E(X_{(1)}X_{(2)}) + 1 \quad ; \quad (2)$$

$$\text{The trace of } V \text{ is } \text{tr}(V) = n - \sum_{i=1}^n m_i^2 \quad . \quad (3)$$

From (2) we obtain

$$V_{12} = E(X_{(1)})^2 - m_1 m_2 - 1 \quad . \quad (4)$$

Of these, (1) is very well known, (2) is given in, e.g., Govindarajulu (1963), and (3) is easily proved as follows:

$$\text{tr}(V) = \sum_i \{X_{(i)} - m_i\}^2 = \sum_i \{X_{(i)}\}^2 - \sum_i m_i^2 = n - \sum_i m_i^2.$$

We shall also need the results, obtained from the symmetry of V :

$$V_{ij} = V_{ji} = V_{rs} = V_{sr}; \quad r = n+1-i, s = n+1-j. \quad (5)$$

Values of m_i have been extensively tabulated; e.g., for $n \leq 20$, to 10 decimal places (d.p.) in Teichroew (1956), and, for all $n \leq 100$ and at intervals for $n \leq 400$, to 5 d.p. in Harter (1961). The sum $\sum_i m_i^2$ is given for $n \leq 100$, to 5 d.p. in Pearson and Hartley (1972, Table 13) and in Owen (1962, p. 154). Ruben (1954) examines the distributions of m_i from a geometric viewpoint; among the results in his paper he gives moments of the extreme order statistic and tabulates the variance of $X_{(1)}$ i.e., V_{11} , for $n \leq 50$, to 8 d.p. Borenus (1966) has extended this tabulation to $n \leq 120$. These exact values are important in obtaining a good approximation for V , since V_{11} is the most inaccurate term in David and Johnson's formulae. LaBrecque (1973) used the David and Johnson technique, and the correct V_{11} , to calculate certain functions of $m' = (m_1, m_2, \dots, m_n)$ and V . In the next section we use V_{11} and the other identities above to give a considerable improvement over the David-Johnson formulae used alone. By normalization of a row we shall mean keeping certain terms fixed and then multiplying the others by a constant, to ensure that the sum of all elements is 1, as required by (1) above.

2. AN APPROXIMATION FOR V

The calculations for V follow the following steps:

- (a) Insert the correct V_{11} from the tables referenced above;
- (b) Insert the correct V_{12} from (4);
- (c) Insert the rest of row 1 using the David and Johnson formulae;

- (d) Keep V_{11} and V_{12} fixed, and normalize row 1. When row 1 has been calculated, fill in column 1, row n and column n from the symmetry relations (5).
- (e) Apart from terms already calculated from steps (a) through (d) (i.e., V_{21} and V_{2n}), calculate row 2 from the David and Johnson formulae, and normalize row 2. Fill in column 2 and row $n-1$ and column $n-1$ from the symmetry of V . Continue with successive rows until all rows are normalized.

These operations make the top left corner of V_{ij} correct, and the rows more accurate than before; but the trace will not satisfy (3). This identity can be used to give further improvement as follows:

- (f) Change V_{22} and its equal, V_{pp} ($p = n-1$), so that (3) is satisfied; then renormalize row 2 with V_{21} , V_{22} , V_{2n} fixed. Fill in symmetric terms, in columns 2 and $n-1$ and row $n-1$.
- (g) Renormalize successively all rows as for row 2; i.e., leave fixed the diagonal term and terms calculated from symmetry relations with previous rows. The entire matrix V will be complete when row $n/2$ is renormalized, for n even, or row $(n-1)/2$, for n odd. In the latter case, (n odd), the middle row will not satisfy (1). The procedure could be iterated to improve this, but our experience suggests that this is not necessary.

3. ACCURACY OF THE METHOD

When the David-Johnson formulae are used alone, by far the greatest error, for those values of n (≤ 20) for which comparisons can be made over the entire matrix V , occurs at V_{11} . For this particular entry we can, of course, extend comparisons to $n = 120$; the error is about 0.00440 at $n = 20$ (about 1.6%) and diminishes very slowly to 0.00395 at $n = 120$ (about 2.2%). In our computations we used the algorithm of Cunningham (1969) to give the inverse of the normal distribution; this will give computational errors much smaller than those in the approximation itself. The very slow decrease lends support to misgivings

expressed by David and Johnson on the convergence properties, for extreme values, of their series. Comparisons of other terms are in the Table; we have selected those terms where either the

TABLE

Comparison of True Values With Two Approximations. The Asterisk Means Maximum Error in the V Matrix for That Approximation.

N	Element	True	Approximation: D-J		D-S	
			Value	Error	Value	Error
10	V_{23}	.146623	.146423	.000200	.146588	.000035*
10	V_{33}	.175003	.174760	.000237*	.174998	.000005
15	V_{22}	.179122	.179271	.000149*	.179090	.000031*
18	V_{13}	.094617	.094546	.000072	.094653	.000036*
18	V_{22}	.166293	.166504	.000211*	.166279	.000014
20	V_{22}	.159573	.159809	.000236*	.159519	.000046*

David-Johnson formulae used alone give largest error (omitting V_{11}) or where the new technique gives largest error; for $n = 15$ and $n = 20$ these both occur at V_{22} . The new method reduces the maximum error to about one-fifth its previous value. From a percentage point of view, the new approximation gives largest percentage error at V_{1n} , where the true covariance is smallest; this maximum percentage error is of the order of 0.15%; the maximum absolute error is generally less than 0.05%. A comparison of the relative sizes of the errors in the Table, and those quoted above for V_{11} shows how important it is to have the exact values for V_{11} to make a good start in approximating V. We have suggested the upper limit $n = 120$ because exact values are known to this point. However, V_{11} approaches zero like $1/(\ln n)$; in fact, asymptotically $V_{11} \ln n$ has limit $\pi^2/12 = 0.82$ (Cramér, 1946, p. 376), though $V_{11} = 0.85/(\ln n)$ gives a more accurate approximation in the region $110 \leq n \leq 120$ (error less than 0.0001). Thus, the use of the algorithm could be extended. In order to

use the David-Johnson formulae, a computer would be needed; the steps given above can be very easily programmed and it seems worthwhile to get the extra accuracy, particularly if, as in some applications, the inverse of V is required. A Fortran program for the entire procedure is available from the authors.

References

- Borenus, G. (1965). On the limit distribution of an extreme value in a sample from a normal distribution. Skand. Aktuarietidskr. 48, 1-15.
- Cramér, H. (1946). Mathematical Methods of Statistics. Princeton: Princeton University Press.
- Cunningham, S. W. (1969). Algorithm AS24: From normal integral to deviate. Appl. Statist. 18, 290-93.
- David, F. N. & Johnson, N. L. (1954). Statistical treatment of censored data. I. Fundamental formulae. Biometrika 41, 228-40.
- Govindarajulu, Z. (1963). On moments of order statistics and quasi-ranges from normal populations. Ann. Math. Statist. 34, 633-51.
- Harter, H. L. (1961). Expected values of normal order statistics. Biometrika 48, 151-65.
- Hastings, C. Jr., Mosteller, F., Tukey, J. W. & Winsor, C. P. (1947). Low moments for small samples: A comparative study of order statistics. Ann. Math. Statist. 18, 413-26.
- La Brecque, J. F. (1973). New goodness-of-fit procedures for the case of normality. Ph.D. thesis, State University of New York at Buffalo.
- Owen, D. B. (1962). Handbook of Statistical Tables. Reading, Mass.: Addison-Wesley.

- Pearson, E. S. & Hartley, H. O. (1972). Biometrika Tables for Statisticians, Vol. 2. Cambridge: Cambridge University Press.
- Plackett, R. L. (1958). Linear estimation from censored data. Ann. Math. Statist. 29, 131-42.
- Ruben, H. (1954). On the moments of order statistics in samples from normal populations. Biometrika 41, 200-27.
- Sarhan, A. E. & Greenberg, B. G. (1956). Estimation of location and scale parameters by order statistics from singly and doubly censored samples. Ann. Math. Statist. 27, 427-51.
- Saw, J. W. (1960). A note on the error after a number of terms of the David-Johnson series for the expected values of normal order statistics. Biometrika 47, 79-86.
- Teichroew, D. (1956). Tables of expected values of order statistics and products of order statistics for samples of size twenty and less from the normal distribution. Ann. Math. Statist. 27, 410-26.

UNCLASSIFIED

SECURITY CLASSIFICATION OF THIS PAGE (When Data Entered)

REPORT DOCUMENTATION PAGE		READ INSTRUCTIONS BEFORE COMPLETING FORM
1. REPORT NUMBER 14	2. GOVT ACCESSION NO.	3. RECIPIENT'S CATALOG NUMBER
4. TITLE (and Subtitle) THE COVARIANCE MATRIX OF NORMAL ORDER STATISTICS		5. TYPE OF REPORT & PERIOD COVERED TECHNICAL REPORT 14
		6. PERFORMING ORG. REPORT NUMBER
7. AUTHOR(s) C.S. Davis and M.A. Stephens		8. CONTRACT OR GRANT NUMBER(s) DAAG29-77-G-0031
9. PERFORMING ORGANIZATION NAME AND ADDRESS Department of Statistics Stanford University Stanford, CA 94305		10. PROGRAM ELEMENT, PROJECT, TASK AREA & WORK UNIT NUMBERS P-14435-M
11. CONTROLLING OFFICE NAME AND ADDRESS U.S. Army Research Office Post Office Box 12211 Research Triangle Park, NC 27709		12. REPORT DATE February 21, 1978
		13. NUMBER OF PAGES 7
14. MONITORING AGENCY NAME & ADDRESS (if different from Controlling Office)		15. SECURITY CLASS. (of this report) UNCLASSIFIED
		15a. DECLASSIFICATION/DOWNGRADING SCHEDULE
16. DISTRIBUTION STATEMENT (of this Report) Approved for Public Release; Distribution Unlimited.		
17. DISTRIBUTION STATEMENT (of the abstract entered in Block 20, if different from Report)		
18. SUPPLEMENTARY NOTES The findings in this report are not to be construed as an official Department of the Army position, unless so designated by other authorized documents. This report partially supported under Office of Naval Research Contract N00014-76-C-0475 (NR-042-267) and issued as Technical Report No. 254.		
19. KEY WORDS (Continue on reverse side if necessary and identify by block number) Normal distribution, covariance matrix calculations, normal order statistics, generalized least squares.		
20. ABSTRACT (Continue on reverse side if necessary and identify by block number) An approximation is given to calculate V , the covariance matrix for normal order statistics. The approximation gives considerable improvement over previous approximations, and the computing algorithm is available from the authors.		

DD FORM 1 JAN 73 1473

EDITION OF 1 NOV 65 IS OBSOLETE
S/N 0102-014-6601

UNCLASSIFIED

SECURITY CLASSIFICATION OF THIS PAGE (When Data Entered)